

Πολυωνυμική $(k-1)$ διάσταση

Πείραμα (ή δοκιμή) με k δυνατά ασυμβίβαστα ενδεχόμενα C_1, C_2, \dots, C_k με πιθανότητα εμφάνισης: $p_i = P(C_i)$, $\sum_{i=1}^k p_i = 1$

n ανεξάρτητες επαναλήψεις του πειράματος.

Εστω τ.μ. X_i = αριθμός εμφάνισης (ή πραγματοποίησης) του C_i στις n επαναλήψεις.
Τότε, $\sum_{i=1}^k X_i = n$ ή n $(k-1)$ διάστατη τ.μ. (X_1, \dots, X_{k-1})

$$P_{X_1, \dots, X_{k-1}}^{(n)}(x_1, \dots, x_{k-1}) = P(X_1=x_1, \dots, X_k=x_k) = \binom{n}{x_1, \dots, x_k} p_1^{x_1} p_2^{x_2} \dots p_k^{x_k} =$$

$$= \frac{n!}{x_1! x_2! \dots x_k!} p_1^{x_1} p_2^{x_2} \dots p_k^{x_k}$$

$$\text{Ετσι: } (X_1, \dots, X_{k-1}) \sim M(n, p_1, p_2, \dots, p_{k-1})$$

Για $k=3$, Τριωνυμική

ΙΔΙΟΤΗΤΕΣ:

$$1) X_i \sim \text{Bin}(n, p_i), \quad i=1, \dots, k-1.$$

$$2) (X_i, X_j) \sim M(n, p_i, p_j), \quad i, j=1, \dots, k-1, \quad i \neq j$$

$$3) X_1 | X_2=x_2, \dots, X_{k-1}=x_{k-1} \sim \text{Bin}\left(n-x_2-\dots-x_{k-1}, \frac{p_1}{1-p_2-\dots-p_{k-1}}\right)$$

$$\text{Για } k=3: \textcircled{1} X_1 \sim \text{Bin}(n, p_1)$$

$$\textcircled{3} X_1 | X_2 \sim \text{Bin}\left(n-x_2, \frac{p_1}{1-p_2}\right)$$

ΑΠΟΔΕΙΞΗ:

$$D) (X_1, X_2) \sim M(n, p_1, p_2) \Rightarrow X_1 \sim \text{Bin}(n, p_1)$$

$$P_{X_1}(x_1) = \sum_{x_2=0}^{n-x_1} \frac{n!}{x_1! x_2! (n-x_1-x_2)!} p_1^{x_1} p_2^{x_2} (1-p_1-p_2)^{n-x_1-x_2} =$$

$$= \frac{n!}{x_1! (n-x_1)!} p_1^{x_1} (1-p_1)^{n-x_1} \sum_{x_2=0}^{n-x_1} \frac{(n-x_1)!}{x_2! (n-x_1-x_2)!} p_2^{x_2} \frac{(1-p_1-p_2)^{n-x_1-x_2}}{(1-p_1)^{n-x_1-x_2} (1-p_1)^{x_2}}$$

$$= \sum_{x_2=0}^{n-x_1} \frac{(n-x_1)!}{x_2! (n-x_1-x_2)!} \left(\frac{p_2}{1-p_1}\right)^{x_2} \underbrace{\left(\frac{1-p_1-p_2}{1-p_1}\right)^{n-x_1-x_2}}_{\left(\frac{1-p_2}{1-p_1}\right)^{x_2}} =$$

$$\equiv \sum_{x_2=0}^{n-x_1} \text{Bin}\left(n-x_1, \frac{p_2}{1-p_1}\right)$$

$$\text{Οποτε τελικά } \equiv \text{Bin}(n, p_1)$$

$$\begin{aligned}
 \text{3) } P_{X_1|X_2}(x_1|x_2) &= \frac{n!}{x_1! x_2! (n-x_1-x_2)!} p_1^{x_1} p_2^{x_2} (1-p_1-p_2)^{n-x_1-x_2} \\
 &= \frac{n!}{x_2! (n-x_2)!} p_2^{x_2} (1-p_2)^{n-x_2} \\
 &= \frac{(n-x_2)!}{x_1! (n-x_1-x_2)!} \cdot \frac{p_1^{x_1} (1-p_1-p_2)^{n-x_1-x_2}}{(1-p_2)^{n-x_1-x_2} (1-p_2)^{x_1}} \\
 &= \frac{(n-x_2)!}{x_1! (n-x_2-x_1)!} \left(\frac{p_1}{1-p_2}\right)^{x_1} \left(1-\frac{p_1}{1-p_2}\right)^{n-x_2-x_1} \\
 &= \text{Bin}(n-x_2, \frac{p_1}{1-p_2})
 \end{aligned}$$

ΠΑΡΑΔΕΙΓΜΑ:

1) $X \equiv$ αριθμός $\xi \cdot 3$, $Y \equiv$ αριθμός $\zeta \cdot 3$, πιθανοί 2 (δύο) Ιαπων.

$n=2$.

$$P_X = \frac{1}{6}, P_Y = \frac{1}{6}$$

$$(X, Y) \sim M(n=2, P_X=1/6, P_Y=1/6)$$

$$X \sim B(n=2, P_X=1/6)$$

$$X|Y=0 \sim \text{Bin}\left(\frac{n-y}{2-0}, p = \frac{P_X}{1-P_Y} = \frac{1/6}{1-1/6} = \frac{1}{5}\right)$$

$$2) X \sim U(9, 15)$$

Επιθυμητή μεταξύ 10 και 12 min.

$$P(X_1=1, X_2=8, X_3=1) = \textcircled{1}$$

$$(X_1, X_2) \sim M(n=10, p_1 = P(X \leq 10) = \int_9^{10} \frac{1}{6} dx = \frac{1}{6}, p_2 = P(10 \leq X \leq 12) = \frac{2}{6})$$

$$X \sim f(x) = \frac{1}{6}, 9 \leq x \leq 15.$$

$$\textcircled{1} = \frac{10!}{1! 8! 1!} \left(\frac{1}{6}\right)^1 \left(\frac{2}{6}\right)^8 \left(\frac{3}{6}\right)^1 = 0.00114$$

Υπεργεωμετρική $(k-1)$ διάταξη $(X_1, \dots, X_{k-1}) \sim Hg(N, n, p_1, \dots, p_{k-1})$
 Πληθυσμός N στοιχεία σε k κατηγορίες C_1, \dots, C_k και έστω Np_i είναι της C_i, \dots, Np_k της C_k με $\sum_{i=1}^k p_i = 1$ δηλ. $\sum_{i=1}^k Np_i = N$

Χωρίς επανάθεση επιλέγονται n στοιχεία στο δείγμα και έστω X_i είναι της $C_i, i=1, \dots, k$
 δηλαδή $\sum_{i=1}^k X_i = n$. Ενδιαφέρει η τ.μ. (X_1, \dots, X_{k-1}) και συμβολίζουμε με:

$$P_{X_1}(\mathcal{X}) = P_{X_1, \dots, X_{k-1}}(X_1, \dots, X_{k-1}) = P(X_1 = x_1, \dots, X_{k-1} = x_{k-1}) = \frac{\binom{Np_1}{x_1} \binom{Np_2}{x_2} \dots \binom{Np_k}{x_k}}{\binom{N}{n}}$$

$$\sum_{i=1}^k p_i = 1, \quad \sum_{i=1}^k x_i = n$$

ΙΔΙΟΤΗΤΕΣ:

- 1) $X_i \sim Hg(N, n, p_i)$
- 2) $(X_1, \dots, X_{k-1}) \sim Hg(N, n, p_1, \dots, p_{k-1})$
 $k_i < k-1$

$$3) X_i | X_2 = x_2, \dots, X_{k-1} = x_{k-1} \sim Hg\left(N - Np_2 - \dots - Np_{k-1}, n - x_2 - \dots - x_{k-1}, \frac{p_1}{1 - p_2 - \dots - p_{k-1}}\right)$$

ΑΠΟΔΕΙΞΗ:

$$1) k=3.$$

$$P_{X_1}(x_1) = \sum_{x_2=0}^{n-x_1} \frac{\binom{Np_1}{x_1} \binom{Np_2}{x_2} \binom{N-Np_1-Np_2}{n-x_1-x_2}}{\binom{N}{n}} = \frac{\binom{Np_1}{x_1}}{\binom{N}{n}} \sum_{x_2=0}^{n-x_1} \binom{Np_2}{x_2} \binom{N-Np_1-Np_2}{n-x_1-x_2}$$

επειδή

$$\sum_{r=0}^k \binom{m}{r} \binom{n}{k-r} = \binom{m+n}{k}$$

$$= \frac{\binom{Np_1}{x_1} \binom{Np_2 + N - Np_1 - Np_2}{x_2 + n - x_1 - x_2}}{\binom{N}{n}}$$

$$= \frac{\binom{Np_1}{x_1} \binom{N - Np_1}{n - x_1}}{\binom{N}{n}} \equiv Hg(N, n, p_1)$$

$$P_{X_1|X_2}(x_1|x_2) = \frac{\binom{Np_1}{x_1} \binom{Np_2}{x_2} \binom{N-Np_1-Np_2}{n-x_1-x_2}}{\binom{Np_2}{x_2} \binom{N-Np_2}{n-x_2}} = \frac{\binom{Np_1}{x_1} \binom{N-Np_1-Np_2}{n-x_1-x_2}}{\binom{N-Np_2}{n-x_2}}$$

$$= \begin{pmatrix} (N-NP_2)P_1 \\ 1-P_2 \\ x_2 \end{pmatrix} \begin{pmatrix} (N-NP_2) \left(1 - \frac{P_1}{1-P_2}\right) \\ n-x_2-x_2 \end{pmatrix} \equiv \text{Hg} \left(N-NP_2, n-x_2, \frac{P_1}{1-P_2} \right)$$

Πολυδιάστατη κανονική

$k=2$

$$f_{X,Y}(x,y) = \frac{1}{2\pi \cdot \sigma_x \sigma_y \sqrt{1-\rho^2}} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x} \right)^2 - 2\rho \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) + \left(\frac{y-\mu_y}{\sigma_y} \right)^2 \right] \right\}$$

$$-1 \leq \rho \leq 1$$

$$-\infty < x, y < \infty, \quad -\infty < \mu_x, \mu_y < \infty, \quad \sigma_x, \sigma_y > 0$$

$$(X, Y) \sim N \left(\begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix}, \begin{pmatrix} \sigma_x^2 & \rho \sigma_x \sigma_y \\ \rho \sigma_x \sigma_y & \sigma_y^2 \end{pmatrix} \right), \quad \rho = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}X \cdot \text{Var}Y}} = \frac{\sigma_{xy}}{\sqrt{\sigma_x^2 \sigma_y^2}} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$

$$\rightarrow \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1 \Rightarrow \frac{x-\mu_x}{\sigma_x} = u \Rightarrow dx = \sigma_x du \text{ και αντίστοιχα } dy = \sigma_y dv.$$

$$\frac{1}{2\pi \cdot \sigma_x \sigma_y \sqrt{1-\rho^2}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2(1-\rho^2)} [u^2 - 2\rho uv + v^2 + \rho^2 v^2 - \rho^2 v^2] \right\} \sigma_x \sigma_y du dv.$$

$$= \frac{1}{2\pi \sqrt{1-\rho^2}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2(1-\rho^2)} [(u-\rho v)^2 + (1-\rho^2)v^2] \right\} du dv =$$

$$= \frac{1}{2\pi \sqrt{1-\rho^2}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2} \left[\frac{u-\rho v}{\sqrt{1-\rho^2}} \right]^2 - \frac{1}{2} v^2 \right\} du dv$$

$$\rightarrow \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{1}{\sqrt{1-\rho^2}} \exp \left\{ -\frac{1}{2} w^2 - \frac{1}{2} v^2 \right\} \sqrt{1-\rho^2} dw dv$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} w^2} dw \cdot \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} v^2} dv = 1 \cdot 1 = 1$$

$$\cdot \frac{u-\rho v}{\sqrt{1-\rho^2}} = w \rightarrow du = \sqrt{1-\rho^2} \cdot dw$$

ΙΔΙΟΤΗΤΕΣ:

$$1) X \sim N(\mu_x, \sigma_x^2), \quad Y \sim N(\mu_y, \sigma_y^2)$$

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy = \frac{1}{2n \sigma_x \sigma_y \sqrt{1-\rho^2}} \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x} \right)^2 - 2\rho \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) + \left(\frac{y-\mu_y}{\sigma_y} \right)^2 \right] \right\} \sigma_y dv =$$

$$= \frac{1}{\sqrt{2n} \cdot \sigma_x} e^{-\frac{1}{2} \left(\frac{x-\mu_x}{\sigma_x} \right)^2} \frac{1}{\sqrt{2n} \sqrt{1-\rho^2}} \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[v^2 - 2\rho v \left(\frac{x-\mu_x}{\sigma_x} \right) + \rho^2 \left(\frac{x-\mu_x}{\sigma_x} \right)^2 \right] \right\} dv$$

$$= N(\mu_x, \sigma_x^2) \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2} w^2 \right\} \sqrt{1-\rho^2} dw = \left\{ \begin{array}{l} v - \rho \left(\frac{x-\mu_x}{\sigma_x} \right) = w \Rightarrow dv = \sqrt{1-\rho^2} dw \\ \sqrt{1-\rho^2} \end{array} \right\}$$

$$= N(\mu_x, \sigma_x^2) \int_{-\infty}^{\infty} f(z) dz = N(\mu_x, \sigma_x^2)$$

$Z \sim N(0, 1)$

$$2) X|Y=y \sim N\left(\mu_x + \rho \frac{\sigma_x}{\sigma_y} (y - \mu_y), (1-\rho^2)\sigma_x^2\right)$$

$$Y|X=x \sim N\left(\mu_y + \rho \frac{\sigma_y}{\sigma_x} (x - \mu_x), (1-\rho^2)\sigma_y^2\right)$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{\frac{1}{2n \sigma_x \sigma_y \sqrt{1-\rho^2}} e^{-\frac{1}{2} \left[\frac{1}{1-\rho^2} \left(\frac{x-\mu_x}{\sigma_x} \right)^2 - \frac{2\rho}{1-\rho^2} \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) + \frac{1}{1-\rho^2} \left(\frac{y-\mu_y}{\sigma_y} \right)^2 \right]}}{\frac{1}{\sqrt{2n} \cdot \sigma_y} e^{-\frac{1}{2} \left(\frac{y-\mu_y}{\sigma_y} \right)^2}}$$

$$= \frac{1}{\sqrt{2n} \cdot \sigma_y \sqrt{1-\rho^2}} e^{-\frac{1}{2} \left[\frac{1}{1-\rho^2} \left(\frac{x-\mu_x}{\sigma_x} \right)^2 - \frac{2\rho}{1-\rho^2} \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) + \frac{\rho^2}{1-\rho^2} \left(\frac{y-\mu_y}{\sigma_y} \right)^2 \right]}$$

$$= \frac{1}{\sqrt{2n} \cdot \sigma_y \sqrt{1-\rho^2}} e^{-\frac{1}{2} \left[\frac{x - \mu_x - \rho \frac{\sigma_x}{\sigma_y} (y - \mu_y)}{\sigma_x \sqrt{1-\rho^2}} \right]^2}$$

$$= \frac{1}{\sqrt{2n} \cdot \sigma_y \sqrt{1-\rho^2}} e^{-\frac{1}{2} \left[\frac{x - \left(\mu_x + \rho \frac{\sigma_x}{\sigma_y} (y - \mu_y) \right)}{\sigma_y \sqrt{1-\rho^2}} \right]^2} \equiv N\left(\mu, \sigma\right)$$

$$3) \underline{X} = (X_1, \dots, X_k) \sim N(\underline{\mu}, \Sigma) \quad \mu \in \mathbb{R}^k \quad f_{\underline{X}}(\underline{x}) = \frac{1}{(\sqrt{2\pi})^k |\Sigma|^{1/2}} e^{-\frac{1}{2} (\underline{x} - \underline{\mu}) \cdot \Sigma^{-1} (\underline{x} - \underline{\mu})}$$

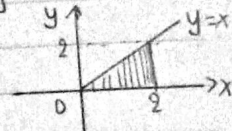
$$\underline{\mu} = \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_k \end{pmatrix} \quad \Sigma = (\sigma_{ij})_{k \times k} = \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

HOMEWORK

ΑΣΚΗΣΗ 23: Έστω X τ.μ. $X \sim U(0,1)$ και έστω η τ.μ. $Y|X=x \sim B(n, p=x)$
 Να βρεθεί η κατανομή της Y .

ΑΣΚΗΣΗ 21: $f_{X,Y}(x,y) = \frac{1}{2}xy I_{(0,x)}(y) I_{(0,2)}(x) = \frac{1}{2}xy$ $0 < x < 2, 0 < y < x$
 $0 < y < x < 2$

Να βρεθούν οι κατανομές των X και Y (ή η $F_{X,Y}(x,y)$)



ΛΥΣΗ:
 $f_X(x) = \int_0^x \frac{1}{2}xy dy = \frac{x^3}{4}, 0 < x < 2.$

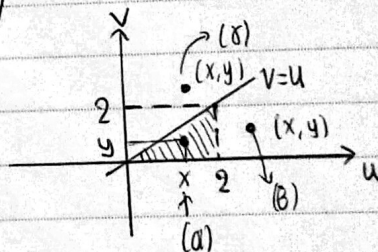
$f_Y(y) = \int_y^2 \frac{1}{2}xy dx = \frac{y(4-y^2)}{4}, 0 < y < 2$

$f_{Y|X}(y|x) = \frac{f(x,y)}{f(x)} = \frac{\frac{1}{2}xy}{\frac{x^3}{4}} = \frac{2y}{x^2}, 0 < y < x < 2.$

$f_{Y|X=1}(y|x=1) = 2y, 0 < y < 1$ ((Beta(a=2, b=1))

$F_{X,Y}(x,y) = 0, x < 0$ ή $y < 0$

$F_{X,Y}(x,y) = 1, x > 2$ και $y > 2$



(a) $= \int_0^y \int_0^x \frac{1}{2}uv du dv = \frac{x^2y^2}{8} - \frac{y^4}{16}, 0 < x < 2$ ή $y < x$

(b) $= \int_0^y \int_2^x \frac{1}{2}uv du dv = \frac{y^2}{2} - \frac{y^4}{16}, x > 2, y < 2.$

(gamma) $= \int_0^x \int_0^u \frac{1}{2}uv dv du = \frac{x^4}{16}, 0 < x, y < 2$ και $y > x$